

Do Small Classes Reduce the Achievement Gap between Low and High
Achievers? Evidence from Project STAR

by

Spyros Konstantopoulos
Northwestern University

Abstract

Given that previous findings on the social distribution of the effects of small classes have been mixed and inconclusive, in the present study I attempted to shed light on the mechanism through which small classes affect the achievement of low- and high-achieving students. I used data from a 4-year large-scale randomized experiment (project STAR) to examine the effects of small classes on the achievement gap. The sample consisted of nearly 11,000 elementary school students who participated in the experiment from kindergarten to grade 3. Meta-analysis and quantile regression methods were employed to examine the effects of small classes on the achievement gap in mathematics and reading (Stanford Achievement Test). The results consistently indicated that higher-achieving students benefited more from being in small classes in early grades than other students. The findings also indicated that although all types of students benefited from being in small classes, reductions in class size did not reduce the achievement gap between low and high achievers.

The effects of class size on student achievement have been of great interest to educational researchers and policy makers the last 2 decades. Reducing class size to boost student achievement is a policy option that has gained considerable attention nationwide. Currently, many states and school districts have enacted or are considering class size reduction with the objective of improving academic achievement.

Studies that used high-quality experimental data have consistently demonstrated the positive effects of small classes on average student achievement for all students (e.g., Finn & Achilles, 1990; Krueger, 1999; Nye, Hedges, & Konstantopoulos, 2000b). Specifically, these studies indicated that the average student achievement in small classes (15 students on average) was significantly higher than in regular classes (22 students on average). These findings suggest that reducing class size is a promising intervention that increases academic achievement on average for all students.

However, it is tempting to imagine class size reduction as an educational intervention that increases academic achievement for all students and reduces the achievement gap between lower- and higher-achieving students by producing larger gains for low achievers. The important question of whether class size reduction can reduce the achievement gap and hence affect the academic achievement of low- and high-achieving students differently has not been fully answered thus far. In the present study we attempted to answer this question by

examining differences in achievement variability between small and regular-size classes using data from a 4-year, large-scale, randomized experiment conducted in Tennessee in the mid 1980s. We also examined differences in achievement between students in small and in regular classes in the upper and lower tails of the achievement distribution. Observed differences in achievement variability along these dimensions would indicate that small classes have varied effects on different groups of students.

Examining Effects of Class Size Reduction on the Achievement Gap

Previous work on the effects of class size has focused exclusively on estimating mean differences in student achievement between small and regular-size classes (Kruger, 1999; Nye, Hedges, & Konstantopoulos, 2000b). However, focusing on average differences of achievement distributions between these types of classes is only one way to evaluate effects of class size. A more complete assessment of the effects would also examine differences in the variability of student achievement between small and regular classes as well as varied effects of small classes in the upper and lower tails of the achievement distribution. Specifically, differences in variability (in a specific outcome) between treatment and control groups in experimental studies provide important evidence about interactions between treatments and individuals' characteristics (see Bryk & Raudenbush, 1988). For example, differences in achievement variability between treatment and control

groups may indicate that a treatment has differential effects on different groups of students; that is, some student groups may benefit more from being exposed to the treatment than others. This notion of interaction between treatments and individual characteristics goes back to the pioneering work of Cronbach and Snow (1977). In this study we followed Cronbach's and Snow's definition about interactions and examined whether different groups of students (such as low and high achievers) benefit more or less from receiving a treatment (being in small classes).

Reducing class size can potentially affect the means as well as the variances of the achievement distributions of small and regular-size classes. That is, class size reduction can also produce differences in the variability of student achievement between the two types of classes. Class size reduction can affect student achievement variability in three ways. The variability of student achievement in small classes may be: (a) less than in regular classes, (b) larger than in regular classes, or (c) similar to that in regular classes. Less variability in student achievement in small classes (than in regular classes) suggests that the achievement gap between lower- and higher-achieving students is smaller in these types of classes. In contrast, larger variability in student achievement in small classes (than in regular classes) implies that the achievement gap between lower- and higher-achieving students is wider in small classes. Similar variability in student achievement in small and in regular classes suggests that the achievement

gap between lower- and higher-achieving students is similar in both types of classes. In addition, differences in achievement variability between small and regular classes may indicate that achievement differences in the middle of the achievement distribution are qualitatively different from achievement differences in the tails. For example, achievement differences between low achievers in small and in regular classes may be significantly smaller or larger than achievement differences between average- or high-achieving students in these classes.

Hypotheses about the Class Size Mechanism

Class size reduction can affect the achievement gap in three ways. First, if high achievers benefit more than low achievers from being in small classes, then one would expect more variability in student achievement in small classes than in regular classes. In this case the achievement distribution in small classes will have a higher mean and a larger standard deviation than the achievement distribution in regular classes. This may also indicate that achievement differences between students in small and in regular classes are larger in the upper tail of the achievement distribution (higher-achieving students) than in the lower tail (lower-achieving students). If this hypothesis were true, then small classes would not close the achievement gap between the two groups of students. Second, if low achievers benefit more from being in small classes, then one would expect smaller variability in achievement in small classes than in regular classes. In this case the

achievement distribution in small classes will have a higher mean and a smaller standard deviation than the achievement distribution in regular classes. This may also indicate that achievement differences between students in small and regular classes are larger at the lower end of the achievement distribution (lower-achieving students) than at the upper end (higher-achieving students). If this hypothesis were true, then small classes would close the achievement gap between low and high achievers because low achievers would benefit more than other students from being in smaller classes. Third, if small and regular classes help higher- and lower-achieving students similarly, then one would expect the variability in achievement in small and regular classes to be comparable. In this case the achievement distribution in small classes will have a higher mean than that in regular classes but a comparable standard deviation. That is, the achievement distribution in small classes is simply shifted to the right by about one-fifth of a standard deviation, which is the average achievement benefit reported in previous studies (e.g., Nye et al., 2000b). If this hypothesis were true, then small classes would have no effect on the achievement gap between low and high achievers because low, average, and high achievers would benefit equally from being in small classes.

There are a variety of ways to investigate these hypotheses. First, one could examine the differential effects of small classes on the achievement of low-achieving, minority, and disadvantaged students. Some recent studies investigated

this issue and found weak evidence that small classes help these types of students more than others (e.g., Nye, Hedges, & Konstantopoulos, 2000a, 2002, 2004a). Alternatively, one could examine differences in the variability of achievement in small and regular classes. Notice that differences in achievement variability between the two types of classes indicate that the treatment has varied effects across different types of students (e.g., high and low achievers). That is, the effects of the treatment may be different in the lower and upper tails of the achievement distribution. In the present study we examined differences in the variability of achievement between small and regular classes as well as differences in the upper and lower tails of the achievement distribution in an attempt to better understand the class size mechanism. We use data from a 4-year, large-scale, randomized experiment conducted in Tennessee in the mid 1980s. This study also addressed issues related to the internal validity of Project STAR such as student switching among different types of classes and more-than-intended variability in actual class size within different types of classes.

Research on Educational Interventions

Prior research has shown that treatments can affect both the mean and the variance of a continuous outcome of interest. Studies evaluating effects of educational interventions have indicated that such programs can change not only average student achievement but the variability in achievement as well. For

example, research on resource allocation in schools has indicated that an important criterion variable for reform programs that aim to equalize school funding is not the average per-pupil expenditure across school districts but the variability of per-pupil expenditure across school districts (Bowles & Levin, 1968; Monk, 1981). Reviews of research where students are assigned to different learning conditions (e.g., tutoring, mastery learning, and conventional) have also reported differences among the three conditions in average achievement as well as in achievement variability (Bloom, 1984). Cross-national comparisons of student achievement have also shown that countries with larger achievement gains in central tendency also had larger gains in achievement variability (Coleman, 1985). Specifically, Japan had not only the largest average achievement gain but the largest gain in achievement variability as well. In addition, significant associations between school size, variation in mathematics course taking, sector (private or public school), and variability in achievement have also been reported using High School and Beyond data (Raudenbush & Bryk, 1987). Nonetheless, in the present study it was not obvious how class size reduction will affect achievement variability. Thus, we predicted that achievement variability in small classes would be different from that in regular classes (a nondirectional hypothesis).

Studies of Class Size

The effects of class size reduction on achievement have been examined empirically via various research designs over the past few decades. Numerous experimental and quasi-experimental studies have investigated the effects of class size on student achievement and have been reviewed by Glass and Smith (1979), Glass, Cahen, Smith, and Filby (1982), Hedges and Stock (1983), and Mosteller, Light, and Sacks (1996). Overall, these reviews have indicated that class size reduction has positive effects on student achievement and that these effects become larger as the class size becomes smaller. Nonetheless, the majority of the studies have been small-scale and short term, and although their results may have high internal validity, the generality of their findings may be limited.

Another line of research has examined the effects of class size reduction via education production function studies (see e.g., Hanushek, 1986). Typically such studies compute the association between class size and achievement, adjusting for important student variables such as race/ethnicity, social class, and previous achievement. The interpretation of the results of these econometric studies has been controversial. Although some reviewers have argued that the effects of class size are small and in many studies statistically insignificant (e.g., Hanushek, 1989), others have contended that the magnitude of the estimates of the mean differences in student achievement is a better way to assess class size effects than statistical significance (e.g., Greenwald, Hedges, & Laine, 1996; Hedges, Laine, & Greenwald, 1994). Although most of these studies were large-

scale and hence their results may have high external validity, their internal validity may be limited because it is not obvious that the association between class size and achievement is causal (that is, class size may be endogenous). For example, it is likely that achievement defines class membership. In addition, omitted-variable bias is possible in these large-scale observational studies, and this can bias estimates of class size effects. Finally, the key independent variable (class size) is typically constructed using school size and number of teachers in the school, and hence it is not an accurate but an aggregate measure of class size.

The Tennessee Class Size Experiment

The Tennessee class size experiment, or Project STAR (Student-Teacher Achievement Ratio), is discussed in detail elsewhere (see, e.g., Krueger, 1999; Nye, Hedges, & Konstantopoulos, 2000b). The experiment involved students in 79 elementary schools in 42 districts in Tennessee. During the first year of the study, within each school, kindergarten students were assigned randomly to classrooms in one of three treatment conditions: smaller classes (with 13 to 17 students), larger classes (with 22 to 26 students), or larger classes with a full-time classroom aide. Teachers were also assigned randomly to classes of different types. The assignments of students to classroom types were maintained through the third grade for students who remained in the study. Some students entered the study in the first grade and subsequent grades and were assigned randomly to

classes at that time. Teachers at each subsequent grade level were also assigned randomly to classes as the experimental cohort passed through the grades. Districts had to agree to participate for 4 years, allow school visits for verification of class sizes, interviewing, and data collection, including extra student testing. They also had to allow research staff to assign pupils and teachers randomly to class types and to maintain the assignment of students to class types from kindergarten through grade 3.

Project STAR has high internal validity because, within each school, students and teachers were assigned randomly to classes of different sizes. In addition, because project STAR is a large-scale randomized experiment that includes a broad range of schools and districts (urban, rural, wealthy, and poor) it likely has higher external validity than smaller-scale studies. Moreover, the study was part of the everyday operation of the schools that participated and hence there is a lower likelihood that novelty effects affected the class size estimates.

Previous Findings from Project STAR

Early analyses of Project STAR data indicated that small classes had positive effects on student achievement (Finn & Achilles, 1990). More recent analyses that considered validity threats (e.g., attrition, switching) also demonstrated that small classes increase student achievement (Krueger, 1999; Nye et al., 2000b). Other analyses have shown long-term positive effects of class

reduction on student performance (Finn, Gerber, Achilles, & Boyd-Zaharias, 2001; Krueger & Whitmore, 2001; Nye, Hedges, & Konstantopoulos, 1999).

Project STAR data have also been used to examine the differential effects of class size on the achievement of low-achieving, minority, and disadvantaged students. An early study reported that class size reduction had larger positive effects for minority students (Finn & Achilles, 1990). These average differences were significant for reading achievement for the first 2 years of the experiment. However, more recent studies that used modern and more appropriate statistical methods could not fully replicate the early findings. For example, Nye et al. (2000a) found weak evidence that class size reduction had larger benefits for minority students. The gain was only observed in reading in one of the model specifications that the researchers examined. The differential effects of small classes for disadvantaged students were statistically insignificant in all specifications. In a subsequent study Nye Hedges, and Konstantopoulos (2002) examined the differential effects of small classes for low-achieving students and found no evidence of additional benefits for these students. However, that study involved students who participated in project STAR for 2 consecutive years, and thus did not include new participants who joined the study the following year. Finally, a more recent study that used follow-up data from Project STAR indicated that being in small classes for 4 years may subsequently decrease the race/ethnic achievement gap in reading in grades 4 to 8 (Nye, Hedges, &

Konstantopoulos, 2004a). Nonetheless, overall there is weak evidence of differential effects of small classes for low-achieving, minority, and disadvantaged students.

In the present study we examined how class size reduction affected the achievement gap between low and high achievers. To determine whether small classes had differential effects on different types of students, we computed differences in achievement variability between small and regular classes, and differences in achievement at the upper and lower tails of the achievement distribution. Given previous findings about the differential effects of small classes, one would expect that differences in achievement variability between smaller and larger classes should be small and insignificant and that the small-class effect for lower achievers would not be as important. This is actually the null hypothesis, which states that the variability in achievement between smaller and larger classes is zero and that all students benefit equally from being in small classes. However, given the results from studies of educational interventions, one would expect that class size reduction may increase the mean and the variability in achievement. This would indicate that small classes have higher variability in achievement than regular classes and that high achievers may benefit more from being in small classes than other students. If that were the case, then class size reduction would not reduce the achievement gap.

Method

Validity of Project STAR

In the STAR experiment, as in all longitudinal large-scale studies, fidelity of implementation was compromised somewhat by three factors. First, there was some switching of students among class types in grades 1, 2, and 3. Second, there was student attrition between kindergarten and grade 3. Third, there was some overlap in the actual sizes among different types of classes because of higher-than-designed variability in sample sizes within classes. The effects of these threats to the validity of the experiment were investigated by other researchers who concluded that the threats did not affect the outcome of the experiment in mean differences in achievement (see Krueger, 1999; Nye et al., 2000b).

To ensure the validity of the experiment, it was also crucial that random assignment effectively eliminated preexisting differences between students and teachers assigned to different classrooms. First, the fact that the randomization of students and teachers to classes was carried out by the consortium of researchers who carried out the experiment, and not by school personnel, enhances its credibility. Second, the effectiveness of the randomization was examined in two recent studies that reported no differences on pre-existing characteristics of students or teachers among the assigned conditions (Krueger, 1999; Nye, Konstantopoulos, & Hedges, 2004b). These results are consistent with what one would expect if randomization were successful. Note that these findings cannot prove that the groups did not differ in unobserved variables. However, confirming

that differences in variables that were measured were not observed makes the probability that there are differences in unobserved variables smaller.

In randomized experiments such as project STAR, participants (e.g., students, teachers) have an equal probability of being assigned to treatment groups (e.g., small classes, regular classes, and regular classes with a classroom aide). This suggests that the students (and teachers) assigned to different class types have similar observed and unobserved characteristics. In turn, this indicates that random assignment is orthogonal to observed and unobserved characteristics. The fact that there is no evidence that randomization was not successful facilitates the causal argument in the present study. That is, when randomization is successful, differences in central tendency and variability in achievement are due entirely to the treatment effect. Hence, the causal argument for differences in average achievement holds also for differences in achievement variability. In Project STAR this suggests that the only source of variance heterogeneity in achievement between smaller and larger classes is the differential effect of the treatment (see Raudenbush & Bryk, 1987).

Statistical Analysis

The first part of the analysis involved differences in achievement variability, and hence the outcome variable was the variability in achievement in each classroom, which we computed following the methods provided by Raudenbush and Bryk (1987). The first step involved the computation of the

within classroom residuals. Because students are nested within classrooms we used a two-level model to compute the student-level residuals. Specifically, the first-level model for student i in classroom j is

$$Y_{ij} = \beta_{0j} + \beta_{1j}FEMALE_{ij} + \beta_{2j}MINORITY_{ij} + \beta_{3j}LOWSES_{ij} + e_{ij},$$

where Y represents mathematics or reading achievement for student i in classroom j , $FEMALE$ is a dummy variable for gender, SES is a dummy variable for free or reduced-price lunch eligibility, $MINORITY$ is a dummy variable for minority group membership (indicating that the student was Black, Hispanic, or Asian), and e is a student- and classroom-specific residual. The idea was to adjust for student characteristics in order to compute the residual variation in achievement in each classroom net of student effects. The second-level model for the classroom specific intercept is

$$\beta_{0j} = \gamma_{00} + \eta_j,$$

where γ_{00} is the average student achievement across all classrooms, and η_j is a classroom-specific random effect. The remaining level one coefficients were treated as fixed at the second level. According to Raudenbush and Bryk (1987), the computation of achievement variability within each classroom involves the level one residuals in each classroom and the degrees of freedom involved in the computation of the achievement variability in each classroom. Namely,

$$\hat{\theta}_j = (\sum e_{ij}^2) / v_j,$$

where θ is the residual achievement variation in class j , e represents the student-specific residuals in class j , and v_j indicates the degrees of freedom with which θ is estimated. In our case

$$v_j = n_j - 1,$$

where n is the number of students in classroom j . Further, Raudenbush and Bryk (1987) recommend the log transformation of θ and provide an unbiased estimator namely

$$d_j = 1/2(\log(\hat{\theta}_j) + v_j^{-1}).$$

The term d_j now represents the residual variability in achievement in classroom j and has a known variance $v_j^{-1}/2$. Because the variance of each classroom-specific outcome is known, and these variances differ among classrooms (heterogeneity of variance), the most appropriate method for analyzing these data is meta-analysis (Konstantopoulos & Hedges, 2004; Raudenbush & Bryk, 2002). Specifically, we ran a two-level (mixed effects) meta-analytic model that is expressed in a single level equation as

$$d_j = \delta_0 + \delta_1 \text{SMALLCLASS}_j + u_j + e_j,$$

where δ_0 is the average variability in achievement across all classrooms, δ_1 is the average difference in achievement variability between small and regular classes that needs to be computed, *SMALLCLASS* is a dichotomous variable (1 if small class and 0 otherwise) that represents random assignment, u_j is a classroom-

specific random effect and e is the usual error term. The most important coefficient is δ_1 , which represents the average difference in achievement variability between small and regular classes. We conducted analyses for mathematics and reading achievement separately for each grade, that is, the analyses were repeated eight times. We also ran models adjusting for possible school effects, because students and teachers were not randomly assigned to schools. All analyses were repeated using fixed-effects models for meta-analysis also. Our sample consisted of 325 classrooms in kindergarten, 337 classrooms in the first grade, 324 classrooms in the second grade, and 326 classrooms in the third grade.

The second part of the analysis involved differences in achievement between students in small and in regular classes at the upper and lower tails of the achievement distribution. Specifically, we used quantile regression to estimate the small-class effect at various points on the achievement distribution (see Bushinsky, 1998; Koenker & Bassett, 1978). We ran quantile regressions for mathematics and reading test scores separately for each grade (k, 1, 2, and 3). In each grade mathematics and reading achievement scores were regressed on small-class assignment (taking the value of 1 if a student was in a small class and 0 otherwise). Gender, race/ethnicity, and lower socioeconomic status (SES) were included as covariates. We examined the small-class effect at the lower tail (e.g.,

10th and 25th), the middle (50th quantile), and the upper tail (e.g., 75th and top 90th) of the achievement distribution.

Results

Small Classes and Achievement Variability

The results reported here involve the Stanford Achievement Test (SAT) reading and mathematics scores collected from kindergarten through grade 3 as part of project STAR. First we computed classroom achievement variability (adjusted for student effects) and then regressed this variability on the small-class binary variable to examine small class effects. Results of this analysis are reported in Table 1.

Insert Table 1 Here

Although all 16 regression estimates were positive, which indicates that achievement variability in small classes was larger than that in regular classes, only four estimates were significantly different from zero. This suggests that only 25% of the mean differences in achievement variability between small and regular classes were statistically significant. In kindergarten the differences in classroom achievement variability were significant at the .05 level in mathematics. In subsequent grades the differences in mathematics classroom achievement variability were not significant. In addition, the magnitude of the coefficients was

smaller in grades 1, 2, and 3. This suggests that, in mathematics, class size differences in variability were observed in the first year of the study.

The results in reading were comparable. Specifically, the coefficients in kindergarten and first grade were larger than those in grades 2 and 3. In grade 1 the differences in classroom achievement variability were statistically significant at the .05 level in reading. In other grades the differences in reading achievement variability were not significant. This indicates that, in reading, class size differences in variability are occurred in the second year of the study. Thus, class size differences in achievement variability were observed for both mathematics and reading mainly during the first 2 years of the study (kindergarten and first grade).

Intention-to-Treat Analysis

As in any large-scale, long-term experiment, the implementation of Project STAR deviated from the experimental design. One limitation was that in grades 1, 2, and 3, students who were assigned initially to a specific type of class in one year switched to other types of classes the next year. For example, in the first grade students who were assigned to regular-size and regular-size-with-an-aide classes were randomized again to receive the other treatment condition. Studies have shown that about 50% of the students assigned to one type of regular class in kindergarten were reassigned to the other type of regular-size class in the

first grade (Krueger, 1999; Nye et al., 2000b). With the exception of student switching between regular and regular classes with aide in first grade, the nonrandom transition rates of students among treatment conditions ranged from two to nine percent across grades (see Nye et al., 2000b). It is noteworthy that the transition rates from regular to small classes were consistently eight to nine percent between grades, whereas transition rates from small to regular classes were much lower (2%-4%).

Because student transitions among types of classes were nonrandom, it is possible that the estimates of the class size effects are biased. Research that examined mean differences in achievement between small and regular classes showed no evidence of bias (Krueger, 1999; Nye et al., 2000b). In the present study we examined whether student switching among different types of classes affected differences in achievement variability between small and regular classes. One way to examine the possible effects of this switching is to estimate effects of the treatment as it was originally assigned the first year a student entered the study. This is equivalent to the intention-to-treat analysis typically used in clinical trials. Suppose a student is assigned to a regular class in kindergarten and switches to a small class in first grade. In the intention-to-treat analysis, this student is assumed to be part of the regular-size class in the first grade, although he or she actually received a different type of treatment in that grade. The idea is that, if the intention-to-treat analysis produces estimates of the treatment effect

that are similar to the estimates obtained from the analysis that defines treatment as it was actually received, switching between classrooms would not compromise the internal validity of the experiment. For each grade (1, 2, and 3) we constructed a new variable that we called “original” assignment as a dichotomous variable taking the value of one if a student was originally assigned to a small class and zero otherwise. Then, we reran the analysis discussed earlier in the analysis section for mathematics and reading for grades 1, 2, and 3.

The results of this analysis are reported in Table 2. The structure of Table 2 is identical to that of Table 1. The results of the intention to treat analysis are qualitatively similar to and consistent with those reported in Table 1. Fifteen out of 16 regression estimates (about 94%) were positive, but only four of the estimates were statistically significant. In kindergarten the differences in classroom variability were statistically significant at the .05 level in mathematics. As in Table 1, the magnitude of the coefficients was smaller in grades 1, 2 and 3. The results for reading were comparable. The coefficients in kindergarten and first grade were larger than those in grades 2 and 3. In first grade the differences in classroom variability were statistically significant at the .05 level in reading. In other grades the differences in reading achievement variability were not significant. Overall these results also indicated that class size differences in achievement variability are observed for both mathematics and reading mainly during the first 2 years of the study (kindergarten and first grade).

Insert Table 2 Here

Actual Class Size and Achievement Variability

Although the experimental design had targeted a certain range of class size for each type of classroom (13 to 17 for smaller classes and 22 to 26 for larger classes), there was more than intended variation in small and regular classes. That is, the actual class size ranged from 11 to 20 for small classes and from 15 to 29 for regular classes (see Table 3). As Table 3 shows, there was a modest overlap between the actual class sizes of the three treatment conditions. This larger-than-intended variability in actual class size for each type of classroom and the modest overlap between small and regular classes may have affected the estimate of the treatment effect. Hence, a more complete analysis would examine the association between actual class size and classroom variability in achievement.

Insert Table 3 Here

To conduct this analysis, one needs to construct actual class size and include it as the main independent variable in the meta-analysis regression. This approach, however, has the disadvantage that, although target class size is

assigned randomly, actual class size is not and may be a result of nonrandom unobserved factors that may also be related to the outcome. That is, any relation between actual class size and achievement variability is not necessarily a causal effect. A common way to overcome this problem is to use random assignment as an instrumental variable (IV) for actual class size (see, e.g., Angrist, Imbens, & Rubin, 1996; Nye et al., 2004b). In the IV regression, actual class size is regressed on random assignment and the predicted values of this regression are used in the meta-analysis regression as the main independent variable. The advantage of this procedure is that it yields estimates of the causal effects of actual class size.

The results of this analysis are reported in Table 4. The structure of Table 4 is identical to that of Tables 1 and 2 and the results are similar to those reported in Tables 1 and 2. Specifically, although all 16 regression estimates were negative as expected, only four estimates were significantly different from zero. In kindergarten the differences in classroom variability were statistically significant at the .05 level in mathematics. Again, the magnitude of the coefficients was smaller in grades 1, 2 and 3. The results for reading were comparable. The coefficients in kindergarten and first grade were larger than those in grades 2 and 3. In first grade the association between class size and classroom achievement variability was statistically significant at the .05 level in reading. In other grades the coefficients were not significant. Overall these results also indicated that class size differences in achievement variability were observed for both mathematics

and reading mainly during the first 2 years of the study (kindergarten and first grade). It should be noted that all analyses were replicated using also fixed-effects models for meta-analysis and the results were similar to those obtained using mixed-effects models.

Insert Table 4 Here

Achievement Differences at the Upper and Lower Tails of the Achievement Distribution

Results from the previous analyses suggested that the small-class effect may not be distributed uniformly across the achievement distribution. In fact, the results indicated that the small-class advantage may be larger at the upper tail of the achievement distribution. If high achievers benefit more from being in small classes, then the small-class advantage at the upper tail must be larger than that at the middle part or at the lower tail of the achievement distribution.

Results of the quantile regression analyses are summarized in Table 5. All estimates are in standard deviation units. As expected, all coefficients in the median (or robust) regression were positive, significantly different from zero, and ranged between one-seventh and one-fourth of a standard deviation. These results were similar to those reported in previous studies that estimated mean differences

(see Nye et al., 2000b). The estimates at the lower tail were also positive but smaller than those at the median or at the upper tail of the achievement distribution. Nonetheless, 50% of the estimates at the tenth quantile (grades 1 and 3) and all estimates at the twenty-fifth quantile were statistically significant. This indicates that lower-achieving students benefited from being in small classes. At the upper tail all estimates (75th and 90th quantiles) were positive and statistically significant. The magnitude of the coefficients indicated that the small class effect was consistently larger for high achievers than for other students. All coefficients estimated in the upper tail of the achievement distribution were much larger than those obtained from the middle or the lower tails. In Kindergarten mathematics the coefficient at the ninetieth quantile was more than twice as large as the coefficient at the fiftieth quantile and nearly four times as large as the coefficient at the tenth quantile. In kindergarten reading the difference in achievement between the median and the ninetieth quantile was much smaller (17%), but the difference between the tenth and the ninetieth quantile estimates was still large. In the first grade, the difference in achievement between the median and the ninetieth quantile was nearly 25% both in mathematics and reading achievement.

Insert Table 5 Here

Although these results seem to support the notion that higher-achieving students may benefit more than other students from being in small classes in the same types of classes, one needs to examine whether the estimates across the different quantiles were statistically significant. Table 6 summarizes t-tests that examined this question. The first column of Table 6 indicates that in grades K and 2 the differences between the small-class effect at the tenth and the ninetieth quantiles were statistically significant at the .05 level in mathematics. Also, in grade 1 differences between the small-class effects at the tenth and the ninetieth quantiles were statistically significant at the .05 level in reading. This indicates that in some grades the very high achievers benefited significantly more from being in small classes than did very low achievers. This finding partly replicates that from the previous analyses that pointed to significant differences in achievement variability in kindergarten in mathematics and grade 1 in reading. The results in column 2 indicate that 50% of the differences between the estimates at the twenty-fifth and the seventy-fifth quantiles were statistically significant. The results in columns 3 and 4 show that some differences between the estimates at the fiftieth and the tenth or ninetieth quantiles were statistically significant. Overall, these results provided some evidence that higher achievers benefited more than other students from being in smaller classes. However, these results do not indicate that low achievers are better off in regular classes, that is, all types of students benefit from being in small classes.

Insert Table 6 Here

Conclusion

Previous work that used Project STAR data provided consistent evidence that being in small classes in early grades leads to higher student achievement on average. Given that class size reduction is an educational intervention that benefits all students by increasing their achievement it is tempting to expect that it could also reduce the achievement gap between higher and lower achievers. However, previous research provided weak or no evidence that reducing class size benefits lower-achieving students more than other students (Nye et al., 2002). The present study examined differences in achievement variability between smaller and larger classes and differences at the upper and lower tails of the achievement distribution in an attempt to better understand the effects of class size reduction on the achievement gap.

Our results suggest that small classes produce significantly higher variability in achievement than regular classes in kindergarten in mathematics and in first grade in reading. The differences favoring small classes were more pronounced and significant in the first 2 years of the experiment (kindergarten and first grade) and smaller and insignificant in the last 2 years of the experiment

(grades 2 and 3). Overall the results indicate that class size reduction increases not only achievement for all students on average, but the variability in student achievement as well (at least in the first two grades). In addition, results from the quantile regression analyses provided additional evidence that all types of students benefited from being in small classes, and that high-achieving students may have benefited even more.

These findings suggest differential effects of small classes across different types of students, that is, some types of students benefit more than others from being in small classes. Specifically, due to the larger variability in achievement in small classes, the difference (or distance) in achievement between high and low achievers is greater in those classes than in regular classes in kindergarten and first grade. If the achievement distributions in small and regular classes had the same mean but different variances (e.g., larger variances in small classes), then one would argue that high achievers may benefit more from being in small classes than in regular classes, whereas low achievers may benefit less. However, the achievement distributions differ in the means as well because smaller classes have a higher mean. This still indicates that higher achievers may benefit more from being small classes, but, given the considerable average difference in achievement (nearly 0.2 SD) low achievers in small classes would benefit at least as much as low achievers in regular classes because the small-class achievement distribution is shifted to the right. Results of the quantile regression analysis support this

notion showing that high achievers benefit even more from being in small classes, but low achievers benefit as well. Nonetheless, the achievement gap between lower and higher achievers is still larger in small classes than in regular classes in some grades. This suggests no evidence that manipulating class size can reduce the achievement gap between lower- and higher achieving students.

Our analyses also addressed the possible effects of validity threats such as student switching between types of classrooms, and larger variability than intended by design in actual class size, which resulted in overlap in actual class size between smaller and larger classes. The results of these analyses were consistent with those in the original analysis and further supported the notion that achievement variability is larger in small classes especially during kindergarten and first grade. This again suggests that high achievers may benefit even more from being small classes than in other types of classes, at least in kindergarten and first grade. However, we did not find any evidence of additional benefits of small classes for lower achievers. This result should be interpreted with caution. It does not necessarily mean that lower-achieving students are better off in larger classes, because all students benefit from being in small classes.

These results shed some more light on the mechanism through which small classes may benefit students. One hypothesis is that in small classes teachers are more likely to identify lower achievers and hence they are more likely to provide instruction designed to benefit these students in the early grades.

However, our findings did not support this hypothesis. Another hypothesis is that teachers are also more likely to identify higher-achieving students in small classes and thus are more likely to provide effective strategies that benefit these students more. Alternatively, it is plausible that the instructional practices in small classes benefit higher achievers more. That is, high-achieving students may be more engaged (or motivated) in learning than other students in small classes. Possibly high achievers take more advantage of the opportunities or teacher practices that take place in small classes or create more opportunities for learning in small classes than lower-achieving students in small classes or other students in other classes, especially in the first 2 years of school (kindergarten and first grade). One possibility is that the effects of small classes accrued mainly in the first and second years of the study. Some researchers has discussed that possibility and showed that the cumulative effects of small classes diminish over time in mathematics (Nye, Hedges, & Konstantopoulos, 2001). Typically in the first 2 years of schooling students learn what behaviors are expected in school. This means that teachers spend considerable time on management and behavior-related issues. It is likely that in smaller classes these issues are addressed in a shorter time than in regular classes, and this in turn means that in kindergarten and first grade more time is spent on learning and instruction in small classes. High achievers in small classes may take advantage of this and engage more in learning than other students. Hence, they may have steeper learning trajectories in the first

2 years of school than other students or in later grades. By grade 2, students typically know what is expected in elementary school and hence it is likely that the time spent on management issues and learning and instruction is comparable in small and regular classes. It is difficult to know exactly what the mechanism might be. Our results, however, indicate that higher-achieving students benefit even more from being in small classes mainly in Kindergarten and first grade.

In addition, it should be noted that the results reported in this study are depended on the outcome used. Specifically, although SAT tests are used widely in elementary schools, it is not obvious that they portray accurately the teacher practices in small and regular classes or the content covered at each grade level. SAT are norm-referenced tests that are not particularly well-aligned with the curriculum taught at each grade level. Suppose for example, that teachers in small classes spend more time to help low achievers master the specified curriculum, but they also spend time helping high achievers develop learning skills beyond the specified curriculum or grade level. Suppose also, that tests such as SAT measure achievement that is not necessarily curriculum specific. Because SAT is not a criterion-referenced test designed to gauge mastery learning, it is possible that it would not provide evidence of the achievement gap due only to the specified, taught curriculum. Hence, it is possible that the achievement gap in terms of the specified curriculum is greatly reduced, but that reduction would not be evident in the results of measures that are not particularly well-aligned with the curriculum

taught. It would be possible then, that a criterion-referenced test would have provided different results.

Although this study helped us better understand the effects of small classes on student achievement, the mechanism is still not clearly defined. Unfortunately, data about practices in different types of classrooms are not available. Such detailed observational data could have unveiled the mechanism of small-class effects via information about instructional processes and interactions between students and teachers. A new randomized experiment with the objective of collecting high-quality observational data in the classrooms would provide invaluable information about the effects of small classes.

References

- Angrist, J., Imbens, G. W., & Rubin, D. B. (1996). Identification of causal effects using instrumental variables. Journal of the American Statistical Association, 91, 444-472.
- Bloom, B. S. (1984). The two sigma problem: The search for methods of group-instruction as effective as one-to-one tutoring. Educational Researcher, 13(6), 4-16.
- Bowles, S., & Levin, H. M. (1968). The determinants of scholastic achievement: An appraisal of some recent evidence. Journal of Human Resources, 3, 3-24.
- Bryk, A. S., & Raudenbush, S. W. (1988). Heterogeneity of variance in experimental studies: A challenge to conventional interpretations. Psychological Bulletin, 104, 396-404.
- Buchinsky, M. (1998). Recent advances in quantile regression models: A practical guideline for empirical research. Journal of Human Resources, 33, 89-126.
- Coleman, J. S. (1985). Schools, families, and children (Nora and Edward Ryerson Lecture). Chicago: University of Chicago.
- Cronbach, L. J., & Snow, R. E. (1977). Aptitudes and instructional methods. New York: Irvington.

- Finn, J D., & Achilles, C. M. (1990). Answers and questions about class size: A statewide experiment. American Educational Research Journal, 27, 557-577.
- Finn, J D., Gerber, S. B., Achilles, C. M., & Boyd-Zaharias, J. (2001). The enduring effects of small classes. Teachers College Record, 103, 145-183.
- Glass, G. V., Cahen, L.S., Smith, M. E., & Filby, N. N. (1982). School class size: Research and policy. Beverly Hills, CA: Sage.
- Glass, G. V., & Smith, M. E. (1979). Meta-analysis of research on class size and achievement. Educational Evaluation and Policy Analysis, 1, 2-16.
- Greenwald, R., Hedges, L. V., & Laine, R. D. (1996). The effects of school Resources on student achievement, Review of Educational Research, 66, 361-396.
- Hanushek, E. A. (1986). The economics of schooling: Production and efficiency in public schools. Journal of Economic Literature, 24, 1141-77.
- Hanushek, E. A. (1989). The impact of differential expenditures on school performance. Educational Researcher, 18, 45-51.
- Hedges, L. V., Laine, R., & Greenwald, R. (1994). Does money matter: A meta-analysis of studies of the effects of differential school inputs on student outcomes. Educational Researcher, 23, 5-14.

- Hedges, L. V., & Stock, W. (1983). The effects of class size: An examination of rival hypotheses. American Educational Research Journal, 20, 63-85.
- Koenker, R., & Bassett, G. (1978). Regression quantiles. *Econometrica*, *46*, 33-50.
- Konstantopoulos, S., & Hedges, L.V. (2004). Meta-analysis. In D. Kaplan (Ed.), Handbook of quantitative methodology for the social sciences (pp. 281-297). New York: Sage.
- Krueger, A. B. (1999). Experimental estimates of education production functions. Quarterly Journal of Economics, 114, 497-532.
- Krueger, A. B., & Whitmore, D. M. (2001). The effect of attending a small class in the early grades on college-test taking and middle school test results: Evidence from project STAR. Economic Journal, 111, 1-28.
- Monk, D. H. (1981). Toward a multilevel perspective on the allocation of educational resources. Review of Educational Research, 51, 215-236.
- Mosteller, F., Light, R. J., & Sachs, J. A. (1983). Sustained inquiry in education: Lessons learned from skill grouping and class size. Harvard Educational Review, 66, 797-842.
- Nye, B., Hedges, L.V., & Konstantopoulos, S. (1999). The long-term effects of small classes: A five-year follow-up of the Tennessee class size experiment. Educational Evaluation and Policy Analysis, 21, 127-142.

- Nye, B., Hedges, L.V., & Konstantopoulos, S. (2000a). Do the disadvantaged benefit from the small classes? Evidence from the Tennessee class size experiment. American Journal of Education, 109, 1-26.
- Nye, B., Hedges, L.V., & Konstantopoulos, S. (2000b). Effects of small classes on academic achievement: The results of the Tennessee class size experiment. American Educational Research Journal, 37, 123-151.
- Nye, B., Hedges, L.V., & Konstantopoulos, S. (2001). Are the effects of small classes cumulative? Evidence from the Tennessee class size experiment. Journal of Educational Research, 94, 336-345.
- Nye, B., Hedges, L.V., & Konstantopoulos, S. (2002). Do low achieving students benefit more from small classes? Evidence from the Tennessee class size experiment. Educational Evaluation and Policy Analysis, 24, 201-217.
- Nye, B., Hedges, L.V., & Konstantopoulos, S. (2004a). Do minorities experience larger lasting benefits from small classes? Evidence from a five-year follow-up of the Tennessee class size experiment. Journal of Educational Research, 98, 94-100.
- Nye, B., Konstantopoulos, S, & Hedges, L.V. (2004b). How Large are Teacher Effects? Educational Evaluation and Policy Analysis, 26, 237-257.
- Raudenbush, S. W., & Bryk, A. S. (1987). Examining correlates of diversity. Journal of Educational Statistics, 12, 241-269.

Raudenbush, S. W., & Bryk, A. S. (2002). Hierarchical linear models. Thousand Oaks, CA: Sage.

Table 1. Mean Differences in Classroom Variability between Small and Regular Classes for Mathematics and Reading

	Mathematics			Reading		
	Coefficient	SE	Two-Tailed P-Value	Coefficient	SE	Two-Tailed P-Value
Kindergarten:						
Small class	.069	.029	.019*	.059	.041	.151
Including school fixed effects	.069	.028	.011*	.058	.037	.113
Grade 1:						
Small class	.033	.028	.236	.069	.029	.017*
Including school fixed effects	.033	.024	.168	.061	.022	.006*
Grade 2:						
Small class	.025	.030	.397	.012	.030	.700
Including school fixed effects	.035	.027	.186	.007	.025	.784
Grade 3:						
Small class	.011	.026	.662	.032	.025	.199
Including school fixed effects	.006	.024	.799	.016	.024	.514

* $p < .05$

Table 2. Mean Differences in Classroom Variability between Small and Regular Classes for Mathematics and Reading:
 Intention to Treat Analysis

	Mathematics			Reading		
	Coefficient	SE	Two-Tailed P-Value	Coefficient	SE	Two-Tailed P-Value
Kindergarten:						
Small class	.069	.029	.019*	.059	.041	.151
Including school fixed effects	.069	.028	.011*	.058	.037	.113
Grade 1:						
Small class	.038	.028	.178	.067	.029	.024*
Including school fixed effects	.037	.025	.136	.063	.022	.005*
Grade 2:						
Small class	.021	.031	.492	.001	.031	.973
Including school fixed effects	.024	.028	.395	-.001	.026	.959
Grade 3:						
Small class	.011	.027	.687	.034	.026	.194
Including school fixed effects	.003	.026	.905	.025	.025	.320

* $p < .05$

Table 3. Distribution of Actual Class Size among Types of Classes and Grades

Class Size	Grade K			Grade 1			Grade 2			Grade 3		
	Small	Regular	Regular/Aide	Small	Regular	Regular/Aide	Small	Regular	Regular/Aide	Small	Regular	Regular/Aide
11										2		
12	8			2			3			2		
13	19			14			16			15		
14	22			18			27			17		
15	23		1	31			32			31		
16	31	1		16	1		29	1		31		1
17	24	4	1	33	1		19			27		
18		1	2	6	2		6			10	1	
19		7	6	3	4	3	1	3	3	5		4
20		6	6	1	10	6		2	1		9	3
21		14	12		18	18		7	11		11	12
22		20	20		27	15		23	21		13	16
23		16	21		19	20		20	21		10	14
24		19	14		16	11		22	25		15	14

Table 3 Continued

Class Size	Grade K			Grade 1			Grade 2			Grade 3		
	Small	Regular	Regular/Aide	Small	Regular	Regular/Aide	Small	Regular	Regular/Aide	Small	Regular	Regular/Aide
25		6	6		7	9		9	15		16	15
26		4	3		5	9		6	7		5	12
27		1	6		2	4		4	1		5	8
28			1		1	2		1			2	6
29					1	2		2	2		2	2
Total	127	99	99	124	114	99	133	100	107	140	89	107
Average	14.96	22.16	22.54	15.52	22.47	23.20	15.16	23.29	23.32	15.53	23.42	23.77

Table 4. Effects of Actual Class Size on Classroom Variability for Mathematics and Reading

	Mathematics			Reading		
	Coefficient	SE	Two-Tailed P-Value	Coefficient	SE	Two-Tailed P-Value
Kindergarten:						
Small class	-.009	.004	.017*	-.008	.006	.160
Including school fixed effects	-.009	.004	.010*	-.008	.005	.125
Grade 1:						
Small class	-.004	.004	.274	-.009	.004	.021*
Including school fixed effects	-.004	.003	.190	-.008	.003	.006*
Grade 2:						
Small class	-.003	.004	.404	-.001	.004	.708
Including school fixed effects	-.004	.003	.193	-.0008	.003	.795
Grade 3:						
Small class	-.001	.003	.683	-.004	.003	.220
Including school fixed effects	-.0007	.003	.817	-.002	.003	.555

* $p < .05$

Table 5. Achievement Differences between Small and Regular Classes at Various
Quantiles for Mathematics and Reading

Mathematics:

Grade:	10th Quantile	25th Quantile	50th Quantile	75th Quantile	90th Quantile
K	.084	.105*	.147*	.178*	.356*
1	.162*	.209*	.255*	.302*	.325*
2	.090	.112*	.157*	.236*	.247*
3	.100*	.126*	.138*	.151*	.201*

Reading:

Grade:

K	.063	.158*	.189*	.252*	.221*
1	.145*	.145*	.236*	.290*	.299*
2	.152*	.152*	.152*	.217*	.261*
3	.078	.104*	.207*	.182*	.207*

* $p < .05$

Table 6. T-tests Indicating Differences in Quantile Regression Estimates

Mathematics:

Grade:	10th Vs 90th Quantile	25th Vs 75th Quantile	10th Vs 50th Quantile	90th Vs 50th Quantile
K	2.854*	1.105	1.038	2.313*
1	1.934	2.151*	2.541*	.913
2	2.169*	2.687*	1.159	1.332
3	1.465	.688	.639	1.161

Reading:

Grade:	10th Vs 90th Quantile	25th Vs 75th Quantile	10th Vs 50th Quantile	90th Vs 50th Quantile
K	1.468	2.243*	2.737*	.368
1	2.651*	3.452*	1.801	1.355
2	1.742	1.264	.000	2.270*
3	1.835	1.465	2.638*	.000

* p < .05